Frontier Topics in Empirical Economics: Week 8 Causal Inference with Panel Data I

Zibin Huang¹

¹College of Business, Shanghai University of Finance and Economics

November 30, 2023

- In the previous lectures, we mostly consider only cross-sectional data
- What if we have one more dimension: Time?
- We call it Panel Data
- We can exploit variations across time for the same individual (unit)

In the previous lectures, we mostly consider only cross-sectional data

- What if we have one more dimension: Time?
- We call it Panel Data
- We can exploit variations across time for the same individual (unit)

- In the previous lectures, we mostly consider only cross-sectional data
- What if we have one more dimension: Time?
- We call it Panel Data
- We can exploit variations across time for the same individual (unit)

- In the previous lectures, we mostly consider only cross-sectional data
- What if we have one more dimension: Time?
- We call it Panel Data
- We can exploit variations across time for the same individual (unit)

- In the previous lectures, we mostly consider only cross-sectional data
- What if we have one more dimension: Time?
- We call it Panel Data
- We can exploit variations across time for the same individual (unit)

- What is the impact of military service on wages?
- Person *i*, Time *t*, Wage Y_i, Military service status D_{it}, Ability A_i, Covariates X_{it}
 Assume constant TE, we have:

$$Y_{it} = \alpha + \rho D_{it} + A'_{i\gamma} + X'_{it}\beta + \epsilon_{it}$$
(1)

$$E[Y_{it}|A_i, X_{it}, D_{it}] = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta$$
⁽²⁾

A_i is the unobserved confounding factor, ε_{it} ⊥ D_{it} |A_i, X_{it}
 How to estimate ρ? Three simple ways

Fixed Effect: FE Settings

What is the impact of military service on wages?

Person *i*, Time *t*, Wage *Y_i*, Military service status *D_{it}*, Ability *A_i*, Covariates *X_{it}*Assume constant TE, we have:

$$Y_{it} = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta + \epsilon_{it}$$
(1)

$$E[Y_{it}|A_i, X_{it}, D_{it}] = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta$$
⁽²⁾

- A_i is the unobserved confounding factor, $\epsilon_{it} \perp D_{it} | A_i, X_{it}$
- How to estimate ρ ? Three simple ways

- What is the impact of military service on wages?
- Person *i*, Time *t*, Wage *Y_i*, Military service status *D_{it}*, Ability *A_i*, Covariates *X_{it}* Assume constant TE, we have:

$$Y_{it} = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta + \epsilon_{it}$$
⁽¹⁾

$$E[Y_{it}|A_i, X_{it}, D_{it}] = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta$$
⁽²⁾

- A_i is the unobserved confounding factor, $\epsilon_{it} \perp D_{it} | A_i, X_{it}$
- How to estimate ρ ? Three simple ways

- What is the impact of military service on wages?
- Person *i*, Time *t*, Wage Y_i , Military service status D_{it} , Ability A_i , Covariates X_{it}
- Assume constant TE, we have:

$$Y_{it} = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta + \epsilon_{it}$$
⁽¹⁾

$$E[Y_{it}|A_i, X_{it}, D_{it}] = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta$$
⁽²⁾

- A_i is the unobserved confounding factor, $\epsilon_{it} \perp D_{it} | A_i, X_{it}$
- How to estimate ρ ? Three simple ways

- What is the impact of military service on wages?
- Person *i*, Time *t*, Wage Y_i , Military service status D_{it} , Ability A_i , Covariates X_{it}
- Assume constant TE, we have:

$$Y_{it} = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta + \epsilon_{it}$$
⁽¹⁾

$$E[Y_{it}|A_i, X_{it}, D_{it}] = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta$$
⁽²⁾

A_i is the unobserved confounding factor, ε_{it} ⊥ D_{it} |A_i, X_{it}
 How to estimate ρ? Three simple ways

- What is the impact of military service on wages?
- Person *i*, Time *t*, Wage Y_i , Military service status D_{it} , Ability A_i , Covariates X_{it}
- Assume constant TE, we have:

$$Y_{it} = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta + \epsilon_{it}$$
⁽¹⁾

$$E[Y_{it}|A_i, X_{it}, D_{it}] = \alpha + \rho D_{it} + A'_i \gamma + X'_{it} \beta$$
⁽²⁾

- A_i is the unobserved confounding factor, $\epsilon_{it} \perp D_{it} | A_i, X_{it}$
- How to estimate ρ ? Three simple ways

Method 1: Fixed Effect Estimator

- FE Estimator is a deviation-from-mean estimator
- Step 1: Take individual-level means of both sides of the regression

$$\bar{Y}_{it} = \alpha + \rho \bar{D}_{it} + A'_i \gamma + \bar{X}'_{it} \beta + \bar{\epsilon}_{it}$$

Step 2: Subtract the mean from the original regression

$$Y_{it} - \bar{Y}_{it} = \alpha - \alpha + \rho (D_{it} - \bar{D}_{it}) + A'_{i\gamma} - A'_{i\gamma} + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(3)
= $\rho (D_{it} - \bar{D}_{it}) + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$ (4)

Unobserved time-invariant A_i is canceled out

ust run regression (4) and get ho

Method 1: Fixed Effect Estimator

FE Estimator is a deviation-from-mean estimator

Step 1: Take individual-level means of both sides of the regression

$$\bar{Y}_{it} = \alpha + \rho \bar{D}_{it} + A_i' \gamma + \bar{X}_{it}' \beta + \bar{\epsilon}_{it}$$

Step 2: Subtract the mean from the original regression

$$Y_{it} - \bar{Y}_{it} = \alpha - \alpha + \rho (D_{it} - \bar{D}_{it}) + A'_{i\gamma} - A'_{i\gamma} + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(3)
= $\rho (D_{it} - \bar{D}_{it}) + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$ (4)

Unobserved time-invariant A_i is canceled out

ust run regression (4) and get ho

Method 1: Fixed Effect Estimator

FE Estimator is a deviation-from-mean estimator

Step 1: Take individual-level means of both sides of the regression

$$\bar{Y}_{it} = \alpha + \rho \bar{D}_{it} + A'_i \gamma + \bar{X}'_{it} \beta + \bar{\epsilon}_{it}$$

Step 2: Subtract the mean from the original regression

$$Y_{it} - \bar{Y}_{it} = \alpha - \alpha + \rho (D_{it} - \bar{D}_{it}) + A'_i \gamma - A'_i \gamma + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(3)
$$= \rho (D_{it} - \bar{D}_{it}) + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(4)

Unobserved time-invariant A_i is canceled out

• Just run regression (4) and get ρ

Method 1: Fixed Effect Estimator

- FE Estimator is a deviation-from-mean estimator
- Step 1: Take individual-level means of both sides of the regression

$$\bar{Y}_{it} = \alpha + \rho \bar{D}_{it} + A'_i \gamma + \bar{X}'_{it} \beta + \bar{\epsilon}_{it}$$

Step 2: Subtract the mean from the original regression

$$Y_{it} - \bar{Y}_{it} = \alpha - \alpha + \rho (D_{it} - \bar{D}_{it}) + A'_i \gamma - A'_i \gamma + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(3)
$$= \rho (D_{it} - \bar{D}_{it}) + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(4)

Unobserved time-invariant A_i is canceled out

• Just run regression (4) and get ρ

Method 1: Fixed Effect Estimator

- FE Estimator is a deviation-from-mean estimator
- Step 1: Take individual-level means of both sides of the regression

$$\bar{Y}_{it} = \alpha + \rho \bar{D}_{it} + A'_i \gamma + \bar{X}'_{it} \beta + \bar{\epsilon}_{it}$$

Step 2: Subtract the mean from the original regression

$$Y_{it} - \bar{Y}_{it} = \alpha - \alpha + \rho(D_{it} - \bar{D}_{it}) + A'_i \gamma - A'_i \gamma + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(3)
= $\rho(D_{it} - \bar{D}_{it}) + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$ (4)

Unobserved time-invariant A_i is canceled out

• Just run regression (4) and get ρ

Method 1: Fixed Effect Estimator

- FE Estimator is a deviation-from-mean estimator
- Step 1: Take individual-level means of both sides of the regression

$$\bar{Y}_{it} = \alpha + \rho \bar{D}_{it} + A'_i \gamma + \bar{X}'_{it} \beta + \bar{\epsilon}_{it}$$

Step 2: Subtract the mean from the original regression

$$Y_{it} - \bar{Y}_{it} = \alpha - \alpha + \rho(D_{it} - \bar{D}_{it}) + A'_i \gamma - A'_i \gamma + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(3)

$$= \rho(D_{it} - \bar{D}_{it}) + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$

$$\tag{4}$$

- Unobserved time-invariant A_i is canceled out
- Just run regression (4) and get ρ

Method 1: Fixed Effect Estimator

- FE Estimator is a deviation-from-mean estimator
- Step 1: Take individual-level means of both sides of the regression

$$\bar{Y}_{it} = \alpha + \rho \bar{D}_{it} + A'_i \gamma + \bar{X}'_{it} \beta + \bar{\epsilon}_{it}$$

Step 2: Subtract the mean from the original regression

$$Y_{it} - \bar{Y}_{it} = \alpha - \alpha + \rho(D_{it} - \bar{D}_{it}) + A'_i \gamma - A'_i \gamma + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$
(3)

$$= \rho(D_{it} - \bar{D}_{it}) + (X'_{it} - \bar{X}'_{it})\beta + (\epsilon_{it} - \bar{\epsilon}_{it})$$

$$\tag{4}$$

- Unobserved time-invariant A_i is canceled out
- Just run regression (4) and get ρ

・ロト・西ト・モト・ 日本

Fixed Effect: Dummy Estimator

Method 2: Dummy Estimator

- We can add a set of individual dummies
- Saturate across the individual dimension

$$Y_{it} = (\alpha + A'_i \gamma) + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$

$$Y_{it} = \alpha_i + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$
(5)

- Unobserved A_i is absorbed in dummy α_i
- Just run regression (5) and get ρ
- Dummy regression is identical to FE regression

Fixed Effect: Dummy Estimator

Method 2: Dummy Estimator

We can add a set of individual dummies

Saturate across the individual dimension

$$Y_{it} = (\alpha + A'_i \gamma) + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$

$$Y_{it} = \alpha_i + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$
(5)

- Just run regression (5) and get ρ
- Dummy regression is identical to FE regression

Fixed Effect: Dummy Estimator

Method 2: Dummy Estimator

We can add a set of individual dummies

Saturate across the individual dimension

$$Y_{it} = (\alpha + A'_i \gamma) + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$

$$Y_{it} = \alpha_i + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$
(5)

- Just run regression (5) and get ρ
- Dummy regression is identical to FE regression

- We can add a set of individual dummies
- Saturate across the individual dimension

$$Y_{it} = (\alpha + A'_i \gamma) + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$

$$Y_{it} = \alpha_i + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$
(5)

- Just run regression (5) and get ρ
- Dummy regression is identical to FE regression

- We can add a set of individual dummies
- Saturate across the individual dimension

$$Y_{it} = (\alpha + A'_i \gamma) + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$

$$Y_{it} = \alpha_i + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$
(5)

- Just run regression (5) and get ρ
- Dummy regression is identical to FE regression

- We can add a set of individual dummies
- Saturate across the individual dimension

$$Y_{it} = (\alpha + A'_i \gamma) + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$

$$Y_{it} = \alpha_i + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$
(5)

- Unobserved A_i is absorbed in dummy α_i
- \blacksquare Just run regression (5) and get ρ
- Dummy regression is identical to FE regression

- We can add a set of individual dummies
- Saturate across the individual dimension

$$Y_{it} = (\alpha + A'_i \gamma) + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$

$$Y_{it} = \alpha_i + \rho D_{it} + X'_{it} \beta + \epsilon_{it}$$
(5)

- Unobserved A_i is absorbed in dummy α_i
- Just run regression (5) and get ρ
- Dummy regression is identical to FE regression

- We can run the regression using differencing (across time) variables
- Assume that $\Delta Y_{it} = Y_{it} Y_{it-1}$ means time difference
- Substracting regression in t by t 1, we have:

$$\Delta Y_{it} = \rho \Delta D_{it} + \Delta X_{it}^{\dagger} \beta + \Delta \epsilon_{it}$$
(6)

Unobserved A_i is canceled out by the differencing

Method 3: FD Estimator

We can run the regression using differencing (across time) variables

• Assume that $\Delta Y_{it} = Y_{it} - Y_{it-1}$ means time difference

■ Substracting regression in t by t - 1, we have:

$$\Delta Y_{it} = \rho \Delta D_{it} + \Delta X'_{it} \beta + \Delta \epsilon_{it}$$
(6)

Unobserved A_i is canceled out by the differencing

Method 3: FD Estimator

- We can run the regression using differencing (across time) variables
- Assume that $\Delta Y_{it} = Y_{it} Y_{it-1}$ means time difference
- Substracting regression in t by t 1, we have:

$$\Delta Y_{it} = \rho \Delta D_{it} + \Delta X_{it}^{\prime} \beta + \Delta \epsilon_{it} \tag{6}$$

Unobserved A_i is canceled out by the differencing

- We can run the regression using differencing (across time) variables
- Assume that $\Delta Y_{it} = Y_{it} Y_{it-1}$ means time difference

• Substracting regression in t by t - 1, we have:

$$\Delta Y_{it} = \rho \Delta D_{it} + \Delta X_{it}^{\prime} \beta + \Delta \epsilon_{it}$$
(6)

Unobserved A_i is canceled out by the differencing

- We can run the regression using differencing (across time) variables
- Assume that $\Delta Y_{it} = Y_{it} Y_{it-1}$ means time difference
- Substracting regression in t by t 1, we have:

$$\Delta Y_{it} = \rho \Delta D_{it} + \Delta X_{it}^{\prime} \beta + \Delta \epsilon_{it}$$
(6)

Unobserved A_i is canceled out by the differencing

- We can run the regression using differencing (across time) variables
- Assume that $\Delta Y_{it} = Y_{it} Y_{it-1}$ means time difference
- Substracting regression in t by t 1, we have:

$$\Delta Y_{it} = \rho \Delta D_{it} + \Delta X'_{it} \beta + \Delta \epsilon_{it}$$
(6)

• Unobserved A_i is canceled out by the differencing

- All of them employ the variations across time for the same person
- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ε_{it} follows random walk, FD is better since difference is now uncorrelated

All of them employ the variations across time for the same person

- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ϵ_{it} follows random walk, FD is better since difference is now uncorrelated

- All of them employ the variations across time for the same person
- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ϵ_{it} follows random walk, FD is better since difference is now uncorrelated

- All of them employ the variations across time for the same person
- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ϵ_{it} follows random walk, FD is better since difference is now uncorrelated
- All of them employ the variations across time for the same person
- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ϵ_{it} follows random walk, FD is better since difference is now uncorrelated

- All of them employ the variations across time for the same person
- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ϵ_{it} follows random walk, FD is better since difference is now uncorrelated

- All of them employ the variations across time for the same person
- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ϵ_{it} follows random walk, FD is better since difference is now uncorrelated

- All of them employ the variations across time for the same person
- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ϵ_{it} follows random walk, FD is better since difference is now uncorrelated

- All of them employ the variations across time for the same person
- FE and Dummy estimators are identical Their point estimations, std errs, and other main statistics are the same
- So we usually call FE and Dummy estimators "FE Model"
- FE and FD are the same in two-period case
- FE and FD are different when T > 2
- When ϵ_{it} are uncorrelated shocks, FE is more efficient than FD
- Since FD will create serial correlation
- But when ϵ_{it} follows random walk, FD is better since difference is now uncorrelated

$$Y_{it} = \rho D_{it} + X_{it}^{\dagger} \beta + \lambda_t + \alpha_i + \epsilon_{it}$$
(7)

- This is called Two-way Fixed Effect Model (TWFE)
- Difference-in-Differences (DID) is a special case of TWFE model
- In DID, usually some policy is implemented at higher level (Province, City...)
- D_{it} is binary (whether individual i at time t is treated by the policy)
- We control for Individual/Province/City level FE and time FE.

$$Y_{it} = \rho D_{it} + X'_{it}\beta + \lambda_t + \alpha_i + \epsilon_{it}$$
⁽⁷⁾

- This is called Two-way Fixed Effect Model (TWFE)
- Difference-in-Differences (DID) is a special case of TWFE model
- In DID, usually some policy is implemented at higher level (Province, City...)
- **D**_{*i*t} is binary (whether individual *i* at time *t* is treated by the policy)
- We control for Individual/Province/City level FE and time FE

$$Y_{it} = \rho D_{it} + X'_{it}\beta + \lambda_t + \alpha_i + \epsilon_{it}$$
⁽⁷⁾

- This is called Two-way Fixed Effect Model (TWFE)
- Difference-in-Differences (DID) is a special case of TWFE model
- In DID, usually some policy is implemented at higher level (Province, City...)
- **D**_{it} is binary (whether individual i at time t is treated by the policy)
- We control for Individual/Province/City level FE and time FE

$$Y_{it} = \rho D_{it} + X'_{it}\beta + \lambda_t + \alpha_i + \epsilon_{it}$$
⁽⁷⁾

- This is called Two-way Fixed Effect Model (TWFE)
- Difference-in-Differences (DID) is a special case of TWFE model
- In DID, usually some policy is implemented at higher level (Province, City...)
- **D**_{it} is binary (whether individual i at time t is treated by the policy)
- We control for Individual/Province/City level FE and time FE

$$Y_{it} = \rho D_{it} + X'_{it}\beta + \lambda_t + \alpha_i + \epsilon_{it}$$
⁽⁷⁾

- This is called Two-way Fixed Effect Model (TWFE)
- Difference-in-Differences (DID) is a special case of TWFE model
- In DID, usually some policy is implemented at higher level (Province, City...)
- **D**_{it} is binary (whether individual i at time t is treated by the policy)
- We control for Individual/Province/City level FE and time FE

$$Y_{it} = \rho D_{it} + X'_{it}\beta + \lambda_t + \alpha_i + \epsilon_{it}$$
⁽⁷⁾

- This is called Two-way Fixed Effect Model (TWFE)
- Difference-in-Differences (DID) is a special case of TWFE model
- In DID, usually some policy is implemented at higher level (Province, City...)
- D_{it} is binary (whether individual *i* at time *t* is treated by the policy)
- We control for Individual/Province/City level FE and time FE

$$Y_{it} = \rho D_{it} + X_{it}' \beta + \lambda_t + \alpha_i + \epsilon_{it}$$
⁽⁷⁾

- This is called Two-way Fixed Effect Model (TWFE)
- Difference-in-Differences (DID) is a special case of TWFE model
- In DID, usually some policy is implemented at higher level (Province, City...)
- D_{it} is binary (whether individual *i* at time *t* is treated by the policy)
- We control for Individual/Province/City level FE and time FE

- Example: Card and Krueger (1994) Effects of minimum wage on employment
- On April 1, 1992, New Jersey raised the state minimum wage
- But in its neighbouring state of Pennsylvania, nothing happened
- Card and Krueger collected employment data in fast food restaurants in NJ and PA in Feb 1992 and Nov 1992

Example: Card and Krueger (1994) Effects of minimum wage on employment

- On April 1, 1992, New Jersey raised the state minimum wage
- But in its neighbouring state of Pennsylvania, nothing happened
- Card and Krueger collected employment data in fast food restaurants in NJ and PA in Feb 1992 and Nov 1992

- Example: Card and Krueger (1994) Effects of minimum wage on employment
- On April 1, 1992, New Jersey raised the state minimum wage
- But in its neighbouring state of Pennsylvania, nothing happened
- Card and Krueger collected employment data in fast food restaurants in NJ and PA in Feb 1992 and Nov 1992

- Example: Card and Krueger (1994) Effects of minimum wage on employment
- On April 1, 1992, New Jersey raised the state minimum wage
- But in its neighbouring state of Pennsylvania, nothing happened
- Card and Krueger collected employment data in fast food restaurants in NJ and PA in Feb 1992 and Nov 1992

- Example: Card and Krueger (1994) Effects of minimum wage on employment
- On April 1, 1992, New Jersey raised the state minimum wage
- But in its neighbouring state of Pennsylvania, nothing happened
- Card and Krueger collected employment data in fast food restaurants in NJ and PA in Feb 1992 and Nov 1992

- For restaurant i in state s at time t, we denote: employment Y_{ist}, minimum wage policy change dummy D_{st}
- In this case, $D_{st} = NJ_sd_t$, if t is after the policy change, $d_t = 1$
- Our target: $E[Y_{1ist} Y_{0ist}|D_{st} = 1]$ (ATT)
- Question: We only observe Y_{1ist} for restaurants in NJ (treated state) after policy
- How would the employment evolve without the policy in NJ?

- For restaurant i in state s at time t, we denote: employment Y_{ist}, minimum wage policy change dummy D_{st}
- In this case, $D_{st} = NJ_sd_t$, if t is after the policy change, $d_t = 1$
- Our target: $E[Y_{1ist} Y_{0ist}|D_{st} = 1]$ (ATT)
- Question: We only observe Y_{1ist} for restaurants in NJ (treated state) after policy
- How would the employment evolve without the policy in NJ?

- For restaurant i in state s at time t, we denote: employment Y_{ist}, minimum wage policy change dummy D_{st}
- In this case, $D_{st} = NJ_sd_t$, if t is after the policy change, $d_t = 1$
- Our target: $E[Y_{1ist} Y_{0ist}|D_{st} = 1]$ (ATT)
- Question: We only observe Y_{1ist} for restaurants in NJ (treated state) after policy
- How would the employment evolve without the policy in NJ?

- For restaurant i in state s at time t, we denote: employment Y_{ist}, minimum wage policy change dummy D_{st}
- In this case, $D_{st} = NJ_sd_t$, if t is after the policy change, $d_t = 1$
- Our target: $E[Y_{1ist} Y_{0ist}|D_{st} = 1]$ (ATT)
- Question: We only observe Y_{1ist} for restaurants in NJ (treated state) after policy
- How would the employment evolve without the policy in NJ?

- For restaurant i in state s at time t, we denote: employment Y_{ist}, minimum wage policy change dummy D_{st}
- In this case, $D_{st} = NJ_sd_t$, if t is after the policy change, $d_t = 1$
- Our target: $E[Y_{1ist} Y_{0ist}|D_{st} = 1]$ (ATT)
- Question: We only observe Y_{1ist} for restaurants in NJ (treated state) after policy

How would the employment evolve without the policy in NJ?

- For restaurant i in state s at time t, we denote: employment Y_{ist}, minimum wage policy change dummy D_{st}
- In this case, $D_{st} = NJ_sd_t$, if t is after the policy change, $d_t = 1$
- Our target: $E[Y_{1ist} Y_{0ist}|D_{st} = 1]$ (ATT)
- Question: We only observe Y_{1ist} for restaurants in NJ (treated state) after policy
- How would the employment evolve without the policy in NJ?

- Let's use restaurants in PA (untreated state) as the control group
- Parallel Trend Assumption: there is no different trend across treated/non-treated states if none of them experienced policy changes

$$E[Y_{0ist}|s,t] = \gamma_s + \lambda_t \tag{8}$$

The no treatment potential outcome Y₀, does not vary across dimension s × t
 No terms like η_{st} in E[Y_{0ist}|s, t]

Let's use restaurants in PA (untreated state) as the control group

Parallel Trend Assumption: there is no different trend across treated/non-treated states if none of them experienced policy changes

$$E[Y_{0ist}|s,t] = \gamma_s + \lambda_t \tag{8}$$

The no treatment potential outcome Y₀, does not vary across dimension s × t
 No terms like η_{st} in E[Y_{0ist}|s, t]

- Let's use restaurants in PA (untreated state) as the control group
- Parallel Trend Assumption: there is no different trend across treated/non-treated states if none of them experienced policy changes

$$E[Y_{0ist}|s,t] = \gamma_s + \lambda_t \tag{8}$$

The no treatment potential outcome Y₀, does not vary across dimension s × t
No terms like η_{st} in E[Y_{0ist}|s, t]

- Let's use restaurants in PA (untreated state) as the control group
- Parallel Trend Assumption: there is no different trend across treated/non-treated states if none of them experienced policy changes

$$E[Y_{0ist}|s,t] = \gamma_s + \lambda_t \tag{8}$$

The no treatment potential outcome Y₀, does not vary across dimension s × t
 No terms like η_{st} in E[Y_{0ist}|s, t]

- Let's use restaurants in PA (untreated state) as the control group
- Parallel Trend Assumption: there is no different trend across treated/non-treated states if none of them experienced policy changes

$$E[Y_{0ist}|s,t] = \gamma_s + \lambda_t \tag{8}$$

The no treatment potential outcome Y₀, does not vary across dimension s × t
No terms like η_{st} in E[Y_{0ist}|s, t]

With the parallel trend assumption, we can identify the policy effect δ by running:

$$Y_{ist} = \gamma_s + \lambda_t + \delta D_{st} + \epsilon_{ist} \tag{9}$$

First difference: For same state, dif across time

$$E[Y_{ist}|s = PA, t = Nov] - E[Y_{ist}|s = PA, t = Feb] = \lambda_{Nov} - \lambda_{Feb}$$
(10)
$$E[Y_{ist}|s = NJ, t = Nov] - E[Y_{ist}|s = NJ, t = Feb] = \lambda_{Nov} - \lambda_{Feb} + \delta$$
(11)

Second difference: Difference in trends across states

$$(11) - (10) = \delta \tag{12}$$

• With the parallel trend assumption, we can identify the policy effect δ by running:

$$Y_{ist} = \gamma_s + \lambda_t + \delta D_{st} + \epsilon_{ist} \tag{9}$$

First difference: For same state, dif across time

$$E[Y_{ist}|s = PA, t = Nov] - E[Y_{ist}|s = PA, t = Feb] = \lambda_{Nov} - \lambda_{Feb}$$
(10)

$$E[Y_{ist}|s = NJ, t = Nov] - E[Y_{ist}|s = NJ, t = Feb] = \lambda_{Nov} - \lambda_{Feb} + \delta$$
(11)

Second difference: Difference in trends across states

$$(11) - (10) = \delta \tag{12}$$

12 / 52

• With the parallel trend assumption, we can identify the policy effect δ by running:

$$Y_{ist} = \gamma_s + \lambda_t + \delta D_{st} + \epsilon_{ist} \tag{9}$$

First difference: For same state, dif across time

$$E[Y_{ist}|s = PA, t = Nov] - E[Y_{ist}|s = PA, t = Feb] = \lambda_{Nov} - \lambda_{Feb}$$
(10)

$$E[Y_{ist}|s = NJ, t = Nov] - E[Y_{ist}|s = NJ, t = Feb] = \lambda_{Nov} - \lambda_{Feb} + \delta$$
(11)

Second difference: Difference in trends across states

$$(11) - (10) = \delta \tag{12}$$

12 / 52

• With the parallel trend assumption, we can identify the policy effect δ by running:

$$Y_{ist} = \gamma_s + \lambda_t + \delta D_{st} + \epsilon_{ist} \tag{9}$$

First difference: For same state, dif across time

$$E[Y_{ist}|s = PA, t = Nov] - E[Y_{ist}|s = PA, t = Feb] = \lambda_{Nov} - \lambda_{Feb}$$
(10)

$$E[Y_{ist}|s = NJ, t = Nov] - E[Y_{ist}|s = NJ, t = Feb] = \lambda_{Nov} - \lambda_{Feb} + \delta$$
(11)

Second difference: Difference in trends across states

$$(11) - (10) = \delta \tag{12}$$

12 / 52

We are taking untreated group as the control!



Figure 5.2.1: Causal effects in the differences-in-differences model

> <回> < 三> < 三> < 三> < 三</p>

DID: Test of Parallel Trend

- After the implementation of the policy at t_0 , we can no longer observe Y_{0i} for the treated group
- Thus, we cannot test parallel trend after t_0
- We test parallel trend before t_0 : Pre-trend test
- There are two simple ways to do that

DID: Test of Parallel Trend

- After the implementation of the policy at t₀, we can no longer observe Y_{0i} for the treated group
- **Thus**, we cannot test parallel trend after t_0
- We test parallel trend before t_0 : Pre-trend test
- There are two simple ways to do that

DID: Test of Parallel Trend

- After the implementation of the policy at t₀, we can no longer observe Y_{0i} for the treated group
- Thus, we cannot test parallel trend after t_0
- We test parallel trend before t_0 : Pre-trend test
- There are two simple ways to do that
- After the implementation of the policy at t₀, we can no longer observe Y_{0i} for the treated group
- Thus, we cannot test parallel trend after t_0
- We test parallel trend before t_0 : Pre-trend test
- There are two simple ways to do that

- After the implementation of the policy at t₀, we can no longer observe Y_{0i} for the treated group
- Thus, we cannot test parallel trend after t_0
- We test parallel trend before t_0 : Pre-trend test
- There are two simple ways to do that

1. Draw the changes in Y across time directly



Figure 5.2.2: Employment in New Jersey and Pennsylvania fast-food restaurants, October 1991 to September 1997 (from Card and Krueger 2000). Vertical lines indicate dates of the original Card and Krueger (1994) survey and the October 1996 federal minimum-wage increase.

1. Draw the changes in Y across time directly

Is this a good pre-trend? (Before the first vertical line)



Figure 5.2.2: Employment in New Jersey and Pennsylvania fast-food restaurants, October 1991 to September 1997 (from Card and Krueger 2000). Vertical lines indicate dates of the original Card and Krueger (1994) survey and the October 1996 federal minimum-wage increase.

- 1. Draw the changes in Y across time directly
 - Is this a good pre-trend? (Before the first vertical line)



Figure 5.2.2: Employment in New Jersey and Pennsylvania fast-food restaurants, October 1991 to September 1997 (from Card and Krueger 2000). Vertical lines indicate dates of the original Card and Krueger (1994) survey and the October 1996 federal minimum-wage increase.

What about this?



Figure 5.2.3. Average rates of grade repetition in second grade for treatment and control schools in Germany (from Pischke 2007). The data span a period before and after a change in term length for students outside of Bavaria.

What about this?



Figure 5.2.3. Average rates of grade repetition in second grade for treatment and control schools in Germany (from Pischke 2007). The data span a period before and after a change in term length for students outside of Bavaria.

2. Event Study Regression

- If we have data from -T to T', and the policy D_{it} is implemented at t = 0
- Let D_s be the dummy of whether in the treated group
- Run the following regression

$$Y_{ist} = \gamma_s + \lambda_t + \sum_{\tau \neq -1} \mathbb{1}(t = \tau) \delta_\tau D_s + \epsilon_{ist}$$
(13)

- δ_τ shows the changes of differences in trends between treated and untreated groups
- In DID, we separate the time into two parts: Before t = 0 and after t = 0 We have D_{st} = D_sd_t = 1 only for treated group after policy implementation
- In event study, we give each time point (year/month) a parameter $\delta_{ au}$

2. Event Study Regression

- If we have data from -T to T', and the policy D_{it} is implemented at t = 0
 Let D_s be the dummy of whether in the treated group
 Due the following correction
 - $Y_{ist} = \gamma_s + \lambda_t + \sum_{\tau \neq -1} 1(t = \tau) \delta_\tau D_s + \epsilon_{ist}$ (13)
- δ_{τ} shows the changes of differences in trends between treated and untreated groups
- In DID, we separate the time into two parts: Before t = 0 and after t = 0
 We have D_{st} = D_sd_t = 1 only for treated group after policy implementation
- lacksquare In event study, we give each time point (year/month) a parameter $\delta_ au$

- 2. Event Study Regression
 - If we have data from -T to T', and the policy D_{it} is implemented at t = 0
 - Let D_s be the dummy of whether in the treated group
 - Run the following regression

$$Y_{ist} = \gamma_s + \lambda_t + \sum_{\tau \neq -1} \mathbb{1}(t = \tau) \delta_\tau D_s + \epsilon_{ist}$$
(13)

- δ_{τ} shows the changes of differences in trends between treated and untreated groups
- In DID, we separate the time into two parts: Before t = 0 and after t = 0We have $D_{st} = D_s d_t = 1$ only for treated group after policy implementation
- In event study, we give each time point (year/month) a parameter $\delta_{ au}$

- 2. Event Study Regression
 - If we have data from -T to T', and the policy D_{it} is implemented at t = 0
 - Let D_s be the dummy of whether in the treated group
 - Run the following regression

$$Y_{ist} = \gamma_s + \lambda_t + \sum_{\tau \neq -1} \mathbb{1}(t = \tau) \delta_\tau D_s + \epsilon_{ist}$$
(13)

- δ_{τ} shows the changes of differences in trends between treated and untreated groups
- In DID, we separate the time into two parts: Before t = 0 and after t = 0We have $D_{st} = D_s d_t = 1$ only for treated group after policy implementation
- \blacksquare In event study, we give each time point (year/month) a parameter δ_{τ}

- 2. Event Study Regression
 - If we have data from -T to T', and the policy D_{it} is implemented at t = 0
 - Let D_s be the dummy of whether in the treated group
 - Run the following regression

$$Y_{ist} = \gamma_s + \lambda_t + \sum_{\tau \neq -1} \mathbb{1}(t = \tau) \delta_\tau D_s + \epsilon_{ist}$$
(13)

- δ_{τ} shows the changes of differences in trends between treated and untreated groups
- In DID, we separate the time into two parts: Before t = 0 and after t = 0We have $D_{st} = D_s d_t = 1$ only for treated group after policy implementation
- In event study, we give each time point (year/month) a parameter $\delta_{ au}$

- 2. Event Study Regression
 - If we have data from -T to T', and the policy D_{it} is implemented at t = 0
 - Let D_s be the dummy of whether in the treated group
 - Run the following regression

$$Y_{ist} = \gamma_s + \lambda_t + \sum_{\tau \neq -1} \mathbb{1}(t = \tau) \delta_\tau D_s + \epsilon_{ist}$$
(13)

- δ_τ shows the changes of differences in trends between treated and untreated groups
- In DID, we separate the time into two parts: Before t = 0 and after t = 0We have $D_{st} = D_s d_t = 1$ only for treated group after policy implementation
- \blacksquare In event study, we give each time point (year/month) a parameter δ_{τ}

- 2. Event Study Regression
 - If we have data from -T to T', and the policy D_{it} is implemented at t = 0
 - Let D_s be the dummy of whether in the treated group
 - Run the following regression

$$Y_{ist} = \gamma_s + \lambda_t + \sum_{\tau \neq -1} \mathbb{1}(t = \tau) \delta_\tau D_s + \epsilon_{ist}$$
(13)

- δ_τ shows the changes of differences in trends between treated and untreated groups
- In DID, we separate the time into two parts: Before t = 0 and after t = 0We have $D_{st} = D_s d_t = 1$ only for treated group after policy implementation
- In event study, we give each time point (year/month) a parameter $\delta_{ au}$

- 2. Event Study Regression
 - If we have data from -T to T', and the policy D_{it} is implemented at t = 0
 - Let D_s be the dummy of whether in the treated group
 - Run the following regression

$$Y_{ist} = \gamma_s + \lambda_t + \sum_{\tau \neq -1} \mathbb{1}(t = \tau) \delta_\tau D_s + \epsilon_{ist}$$
(13)

- δ_τ shows the changes of differences in trends between treated and untreated groups
- In DID, we separate the time into two parts: Before t = 0 and after t = 0We have $D_{st} = D_s d_t = 1$ only for treated group after policy implementation
- In event study, we give each time point (year/month) a parameter $\delta_{ au}$

Usually t = -1 (just before the policy) is omitted as the baseline
 Then we draw the changes of δ for each time period and have:



Points before t = 0 is not significant ⇒ Pre-trend is parallel
 Points after t = 0 shows the policy effect

• Usually t = -1 (just before the policy) is omitted as the baseline

Then we draw the changes of δ for each time period and have:



Points before t = 0 is not significant \Rightarrow Pre-trend is parallel

Points after t = 0 shows the policy effect

- Usually t = -1 (just before the policy) is omitted as the baseline
- Then we draw the changes of δ for each time period and have:



- Points before t = 0 is not significant \Rightarrow Pre-trend is parallel
- Points after t = 0 shows the policy effect

- Usually t = -1 (just before the policy) is omitted as the baseline
- Then we draw the changes of δ for each time period and have:



• Points before t = 0 is not significant \Rightarrow Pre-trend is parallel

Points after t = 0 shows the policy effect

- Usually t = -1 (just before the policy) is omitted as the baseline
- Then we draw the changes of δ for each time period and have:



- Points before t = 0 is not significant \Rightarrow Pre-trend is parallel
- Points after t = 0 shows the policy effect

- 1. Draw changes of Y as a descriptive evidence
- 2. Run your main DID regression
- 3. Run event study regression to check the pre-trend and the dynamic effect
- 4. Remember to cluster your standard errors (More details in the following lectures)

• 1. Draw changes of Y as a descriptive evidence

- 2. Run your main DID regression
- **3**. Run event study regression to check the pre-trend and the dynamic effect
- 4. Remember to cluster your standard errors (More details in the following lectures)

- 1. Draw changes of Y as a descriptive evidence
- 2. Run your main DID regression
- **3**. Run event study regression to check the pre-trend and the dynamic effect
- 4. Remember to cluster your standard errors (More details in the following lectures)

- 1. Draw changes of Y as a descriptive evidence
- 2. Run your main DID regression
- 3. Run event study regression to check the pre-trend and the dynamic effect
- 4. Remember to cluster your standard errors (More details in the following lectures)

- 1. Draw changes of Y as a descriptive evidence
- 2. Run your main DID regression
- 3. Run event study regression to check the pre-trend and the dynamic effect
- 4. Remember to cluster your standard errors (More details in the following lectures)

- When you are doing causal research, a central question is: What kind of variations are used to identify the causal effect?
- It is very very very important!!!
- It determines how you can interpret your results
- It determines which assumption you are using
- Then also determines how you should defend your research (why your assumption is reasonable)

- When you are doing causal research, a central question is: What kind of variations are used to identify the causal effect?
- It is very very very important!!!
- It determines how you can interpret your results
- It determines which assumption you are using
- Then also determines how you should defend your research (why your assumption is reasonable)

- When you are doing causal research, a central question is: What kind of variations are used to identify the causal effect?
- It is very very very important!!!
- It determines how you can interpret your results
- It determines which assumption you are using
- Then also determines how you should defend your research (why your assumption is reasonable)

- When you are doing causal research, a central question is: What kind of variations are used to identify the causal effect?
- It is very very very important!!!
- It determines how you can interpret your results
- It determines which assumption you are using
- Then also determines how you should defend your research (why your assumption is reasonable)

- When you are doing causal research, a central question is: What kind of variations are used to identify the causal effect?
- It is very very very important!!!
- It determines how you can interpret your results
- It determines which assumption you are using
- Then also determines how you should defend your research (why your assumption is reasonable)

- When you are doing causal research, a central question is: What kind of variations are used to identify the causal effect?
- It is very very very important!!!
- It determines how you can interpret your results
- It determines which assumption you are using
- Then also determines how you should defend your research (why your assumption is reasonable)

イロト 不得 トイヨト イヨト ヨヨ ののの

- It becomes complicated in panel data ← more dimensions
- Sometimes, people control many FEs at different levels
- Some are even combined with IV, RD, or other regression structure
- Still, you should always be very clear about your identification source

■ It becomes complicated in panel data ← more dimensions

- Sometimes, people control many FEs at different levels
- Some are even combined with IV, RD, or other regression structure
- Still, you should always be very clear about your identification source

- It becomes complicated in panel data ← more dimensions
- Sometimes, people control many FEs at different levels
- Some are even combined with IV, RD, or other regression structure
- Still, you should always be very clear about your identification source

- It becomes complicated in panel data \leftarrow more dimensions
- Sometimes, people control many FEs at different levels
- Some are even combined with IV, RD, or other regression structure
- Still, you should always be very clear about your identification source

- It becomes complicated in panel data ← more dimensions
- Sometimes, people control many FEs at different levels
- Some are even combined with IV, RD, or other regression structure
- Still, you should always be very clear about your identification source
Let's consider a simple case: effects of working experience on wage
 For individual *i* from family *j* at time *t*:

$$wage_{ijt} = \beta_0 + \beta_1 exp_{ijt} + \epsilon_{ijt}$$
(14)

Let's consider a simple case: effects of working experience on wage For individual *i* from family *j* at time *t*:

$$wage_{ijt} = \beta_0 + \beta_1 e x p_{ijt} + \epsilon_{ijt}$$
(14)

- Let's consider a simple case: effects of working experience on wage
- For individual *i* from family *j* at time *t*:

$$wage_{ijt} = \beta_0 + \beta_1 exp_{ijt} + \epsilon_{ijt}$$
(14)

メロト メポト メヨト メヨト 正正 ろくつ

- When controlling for time FE, you are using variations across individuals and families (*i*, *j* level) in the same year
- When controlling for individual FE, you are using variations across time (t level) for the same people
- When controlling for family FE and time FE, you are using variations across individuals within the same family (*i*|*j* level)
- When controlling for individual FE and time FE, you are using variations of time trends for different people ($i \times t$ level)

- When controlling for time FE, you are using variations across individuals and families (*i*, *j* level) in the same year
- When controlling for individual FE, you are using variations across time (*t* level) for the same people
- When controlling for family FE and time FE, you are using variations across individuals within the same family (*i*|*j* level)
- When controlling for individual FE and time FE, you are using variations of time trends for different people (*i* × *t* level)

- When controlling for time FE, you are using variations across individuals and families (*i*, *j* level) in the same year
- When controlling for individual FE, you are using variations across time (t level) for the same people
- When controlling for family FE and time FE, you are using variations across individuals within the same family (*i*|*j* level)
- When controlling for individual FE and time FE, you are using variations of time trends for different people (*i* × *t* level)

- When controlling for time FE, you are using variations across individuals and families (*i*, *j* level) in the same year
- When controlling for individual FE, you are using variations across time (t level) for the same people
- When controlling for family FE and time FE, you are using variations across individuals within the same family (*i*|*j* level)
- When controlling for individual FE and time FE, you are using variations of time trends for different people (*i* × *t* level)

- When controlling for time FE, you are using variations across individuals and families (*i*, *j* level) in the same year
- When controlling for individual FE, you are using variations across time (t level) for the same people
- When controlling for family FE and time FE, you are using variations across individuals within the same family (*i*|*j* level)
- When controlling for individual FE and time FE, you are using variations of time trends for different people (i × t level)

- We have introduced basic methods of causal inference in panel dataNow we go to three important extensions
 - Recent development in pre-trend testing
 - Synthetic Control Method: When you do not have parallel trend
 - a Staggered DID: When policy implementation scheme is complicated

• We have introduced basic methods of causal inference in panel data

- Now we go to three important extensions
 - Recent development in pre-trend testing
 - Synthetic Control Method: When you do not have parallel trenc
 - Staggered DID: When policy implementation scheme is complicated

• We have introduced basic methods of causal inference in panel data

Now we go to three important extensions

- Recent development in pre-trend testing
- Synthetic Control Method: When you do not have parallel trend
- Staggered DID: When policy implementation scheme is complicated

- We have introduced basic methods of causal inference in panel data
- Now we go to three important extensions
 - Recent development in pre-trend testing
 - Synthetic Control Method: When you do not have parallel trend
 - Staggered DID: When policy implementation scheme is complicated

- We have introduced basic methods of causal inference in panel data
- Now we go to three important extensions
 - Recent development in pre-trend testing
 - Synthetic Control Method: When you do not have parallel trend
 - Staggered DID: When policy implementation scheme is complicated

- We have introduced basic methods of causal inference in panel data
- Now we go to three important extensions
 - Recent development in pre-trend testing
 - Synthetic Control Method: When you do not have parallel trend
 - Staggered DID: When policy implementation scheme is complicated

- Is event study a perfect tool to test parallel pre-trend?
- It's good, but far from perfect
- Roth (2022) Pre-test with Caution: Event-study Estimates After Testing for Parallel Trends
- It also has a latest published version in AER Insights

Is event study a perfect tool to test parallel pre-trend?

- It's good, but far from perfect
- Roth (2022) Pre-test with Caution: Event-study Estimates After Testing for Parallel Trends
- It also has a latest published version in AER Insights

- Is event study a perfect tool to test parallel pre-trend?
- It's good, but far from perfect
- Roth (2022) Pre-test with Caution: Event-study Estimates After Testing for Parallel Trends
- It also has a latest published version in AER Insights

- Is event study a perfect tool to test parallel pre-trend?
- It's good, but far from perfect
- Roth (2022) Pre-test with Caution: Event-study Estimates After Testing for Parallel Trends
- It also has a latest published version in AER Insights

- Is event study a perfect tool to test parallel pre-trend?
- It's good, but far from perfect
- Roth (2022) Pre-test with Caution: Event-study Estimates After Testing for Parallel Trends
- It also has a latest published version in AER Insights

1. Statistical power is low: Likely to have type-II error

- Pre-existing trends that produce meaningful bias may not be detected
- Assuming a linear violation of parallel trend: $\delta_{1t} \delta_{0t} = \gamma t$
- Roth implements some Monte Carlo Simulation using data from 70 papers
- He finds that if you want to detect this violation for 80% of the time, the bias has to be as large as the estimated TE! (100% bias)
- The bias has to be very large for you to detect it

1. Statistical power is low: Likely to have type-II error

- Pre-existing trends that produce meaningful bias may not be detected.
- Assuming a linear violation of parallel trend: $\delta_{1t} \delta_{0t} = \gamma t$
- Roth implements some Monte Carlo Simulation using data from 70 papers.
- He finds that if you want to detect this violation for 80% of the time, the bias has to be as large as the estimated TE! (100% bias)
- The bias has to be very large for you to detect it!

- 1. Statistical power is low: Likely to have type-II error
 - Pre-existing trends that produce meaningful bias may not be detected
 - Assuming a linear violation of parallel trend: $\delta_{1t} \delta_{0t} = \gamma t$
 - Roth implements some Monte Carlo Simulation using data from 70 papers
 - He finds that if you want to detect this violation for 80% of the time, the bias has to be as large as the estimated TE! (100% bias)
 - The bias has to be very large for you to detect it!

- 1. Statistical power is low: Likely to have type-II error
 - Pre-existing trends that produce meaningful bias may not be detected
 - Assuming a linear violation of parallel trend: $\delta_{1t} \delta_{0t} = \gamma t$
 - Roth implements some Monte Carlo Simulation using data from 70 papers
 - He finds that if you want to detect this violation for 80% of the time, the bias has to be as large as the estimated TE! (100% bias)
 - The bias has to be very large for you to detect it!

- 1. Statistical power is low: Likely to have type-II error
 - Pre-existing trends that produce meaningful bias may not be detected
 - Assuming a linear violation of parallel trend: $\delta_{1t} \delta_{0t} = \gamma t$
 - Roth implements some Monte Carlo Simulation using data from 70 papers
 - He finds that if you want to detect this violation for 80% of the time, the bias has to be as large as the estimated TE! (100% bias)
 - The bias has to be very large for you to detect it!

- 1. Statistical power is low: Likely to have type-II error
 - Pre-existing trends that produce meaningful bias may not be detected
 - Assuming a linear violation of parallel trend: $\delta_{1t} \delta_{0t} = \gamma t$
 - Roth implements some Monte Carlo Simulation using data from 70 papers
 - He finds that if you want to detect this violation for 80% of the time, the bias has to be as large as the estimated TE! (100% bias)
 - The bias has to be very large for you to detect it!

- 1. Statistical power is low: Likely to have type-II error
 - Pre-existing trends that produce meaningful bias may not be detected
 - Assuming a linear violation of parallel trend: $\delta_{1t} \delta_{0t} = \gamma t$
 - Roth implements some Monte Carlo Simulation using data from 70 papers
 - He finds that if you want to detect this violation for 80% of the time, the bias has to be as large as the estimated TE! (100% bias)
 - The bias has to be very large for you to detect it!

- **Type I error:** H0 is true but we reject it, α
- \blacksquare Type II error: H0 is false but we do not reject it, β
- Significance level: Probability of commiting Type I error, α
- **Power:** Probability of rejecting H0 if it is false, 1β

Type I and Type II Error		
Null hypothesis is	True	False
Rejected	Type I error False positive Probability = α	
Not rejected		Type II error False negative Probability = β
🛇 Scribbr		

Why? It goes back to the nature of the statistical test

Type I error: H0 is true but we reject it, α

- **Type II error: H0 is false but we do not reject it,** β
- Significance level: Probability of commiting Type I error, α
- **Power:** Probability of rejecting H0 if it is false, 1β



- \blacksquare Type I error: H0 is true but we reject it, α
- \blacksquare Type II error: H0 is false but we do not reject it, β
- Significance level: Probability of commiting Type I error, α
- Power: Probability of rejecting H0 if it is false, 1β



- **Type I error**: H0 is true but we reject it, α
- \blacksquare Type II error: H0 is false but we do not reject it, β
- Significance level: Probability of commiting Type I error, α
- Power: Probability of rejecting H0 if it is false, 1β



- \blacksquare Type I error: H0 is true but we reject it, α
- **Type II error**: H0 is false but we do not reject it, β
- Significance level: Probability of commiting Type I error, α
- Power: Probability of rejecting H0 if it is false, 1β



- **Type I error**: H0 is true but we reject it, α
- \blacksquare Type II error: H0 is false but we do not reject it, β
- Significance level: Probability of commiting Type I error, α
- Power: Probability of rejecting H0 if it is false, 1β



Tradeoff!!!

Now you have to choose a threshold critical value to make your rejection decision
 Go left, you have larger α; Go right, you have larger β



Tradeoff!!!

Now you have to choose a threshold critical value to make your rejection decision
 Go left, you have larger α; Go right, you have larger β



Tradeoff!!!

- Now you have to choose a threshold critical value to make your rejection decision
- Go left, you have larger α ; Go right, you have larger β



Tradeoff!!!

- Now you have to choose a threshold critical value to make your rejection decision
- **G** Go left, you have larger α ; Go right, you have larger β


Tradeoff!!!

- You can decrease T1ER by decreasing $\alpha \Rightarrow$ increasing β
- You can decrease T2ER by decreasing $\beta \Rightarrow$ increasing α
- If you want H_0 to be rejected less easily, you have to tolerate large probability to have false negative



Tradeoff!!!

- You can decrease T1ER by decreasing $\alpha \Rightarrow$ increasing β
- You can decrease T2ER by decreasing $\beta \Rightarrow$ increasing α
- If you want H_0 to be rejected less easily, you have to tolerate large probability to have false negative



Tradeoff!!!

• You can decrease T1ER by decreasing $\alpha \Rightarrow$ increasing β

- You can decrease T2ER by decreasing $\beta \Rightarrow$ increasing α
- If you want H_0 to be rejected less easily, you have to tolerate large probability to have false negative



Tradeoff!!!

- You can decrease T1ER by decreasing $\alpha \Rightarrow$ increasing β
- You can decrease T2ER by decreasing $\beta \Rightarrow$ increasing α
- If you want H_0 to be rejected less easily, you have to tolerate large probability to have false negative



Tradeoff!!!

- You can decrease T1ER by decreasing $\alpha \Rightarrow$ increasing β
- You can decrease T2ER by decreasing $\beta \Rightarrow$ increasing α
- If you want H_0 to be rejected less easily, you have to tolerate large probability to have false negative



- In traditional testing, we try to be conservative about rejecting H0
- **•** Minimize Type I error probability α to be smaller than some level (10%, 5%, 1%)
- It then leads to large $\beta! \Rightarrow$ small power
- But in pre-trend testing, actually we care more about power
- We want to be more conservative about NOT rejecting H0

In traditional testing, we try to be conservative about rejecting H0

- Minimize Type I error probability α to be smaller than some level (10%, 5%, 1%)
- It then leads to large $\beta! \Rightarrow$ small power
- But in pre-trend testing, actually we care more about power
- We want to be more conservative about NOT rejecting H0

- In traditional testing, we try to be conservative about rejecting H0
- Minimize Type I error probability α to be smaller than some level (10%, 5%, 1%)
- It then leads to large $\beta! \Rightarrow$ small power
- But in pre-trend testing, actually we care more about power
- We want to be more conservative about NOT rejecting H0

- In traditional testing, we try to be conservative about rejecting H0
- Minimize Type I error probability α to be smaller than some level (10%, 5%, 1%)
- It then leads to large $\beta! \Rightarrow$ small power
- But in pre-trend testing, actually we care more about power
- We want to be more conservative about NOT rejecting H0

- In traditional testing, we try to be conservative about rejecting H0
- Minimize Type I error probability α to be smaller than some level (10%, 5%, 1%)
- It then leads to large $\beta! \Rightarrow$ small power
- But in pre-trend testing, actually we care more about power
- We want to be more conservative about NOT rejecting H0

- In traditional testing, we try to be conservative about rejecting H0
- Minimize Type I error probability α to be smaller than some level (10%, 5%, 1%)
- It then leads to large $\beta! \Rightarrow$ small power
- But in pre-trend testing, actually we care more about power
- We want to be more conservative about NOT rejecting H0

By selecting samples that can pass the test

- 2a. Underestimate the variance of the estimation
- 2b. If there is bias, conditioning on passing the event study test may exacerbate it (Adding bias to point estimation)
 - The bias is certainly exacerbated in common cases (monotone trends and homoskedastic errors)
 - Thus, the effect of pre-trend testing can be ambiguous.
 Boings (pre-preading series) to prevent the prevent of the

By selecting samples that can pass the test

- 2a. Underestimate the variance of the estimation
- 2b. If there is bias, conditioning on passing the event study test may exacerbate it (Adding bias to point estimation)
 - The bias is certainly exacerbated in common cases (monotone trends and homoskedastic errors)
 - Thus, the effect of pre-trend testing can be ambiguous
 Reject non-parallel cases (good) vs. Increasing bias if there is bias (based)

By selecting samples that can pass the test

- 2a. Underestimate the variance of the estimation
- 2b. If there is bias, conditioning on passing the event study test may exacerbate it (Adding bias to point estimation)
 - The bias is certainly exacerbated in common cases (monotone trends and homoskedastic errors)
 - Thus, the effect of pre-trend testing can be ambiguous Reject non-parallel cases (good) vs. Increasing bias if there is bias (bad)

By selecting samples that can pass the test

- 2a. Underestimate the variance of the estimation
- 2b. If there is bias, conditioning on passing the event study test may exacerbate it (Adding bias to point estimation)
 - The bias is certainly exacerbated in common cases (monotone trends and homoskedastic errors)
 - Thus, the effect of pre-trend testing can be ambiguous
 Reject non-parallel cases (good) vs. Increasing bias if there is bias (bad)

By selecting samples that can pass the test

- 2a. Underestimate the variance of the estimation
- 2b. If there is bias, conditioning on passing the event study test may exacerbate it (Adding bias to point estimation)
 - The bias is certainly exacerbated in common cases (monotone trends and homoskedastic errors)
 - Thus, the effect of pre-trend testing can be ambiguous
 Reject non-parallel cases (good) vs. Increasing bias if there is bias (bad)

By selecting samples that can pass the test

- 2a. Underestimate the variance of the estimation
- 2b. If there is bias, conditioning on passing the event study test may exacerbate it (Adding bias to point estimation)
 - The bias is certainly exacerbated in common cases (monotone trends and homoskedastic errors)
 - Thus, the effect of pre-trend testing can be ambiguous
 Reject non-parallel cases (good) vs. Increasing bias if there is bias (bad)

Practical suggestions proposed by Roth

- Most important advice:
 - Always use your economic knowledge to verify the parallel trend assumption!
- Do not think you are safe when event study does not detect anything
- Do power calculations against economically relevant violations of parallel trends (R package, but not Stata...)
- If you know the functional form of the differences in trends, control it
- Sensitivity analysis using Rambachan and Roth (2023)
 Calculate the bounds of your estimates if there is some violation

Practical suggestions proposed by Roth

- Most important advice:
- Always use your economic knowledge to verify the parallel trend assumption!
- Do not think you are safe when event study does not detect anything
- Do power calculations against economically relevant violations of parallel trends (R package, but not Stata...)
- If you know the functional form of the differences in trends, control it
- Sensitivity analysis using Rambachan and Roth (2023)
 Calculate the bounds of your estimates if there is some violation

Practical suggestions proposed by Roth

Most important advice:

- Do not think you are safe when event study does not detect anything
- Do power calculations against economically relevant violations of parallel trends (R package, but not Stata...)
- If you know the functional form of the differences in trends, control it
- Sensitivity analysis using Rambachan and Roth (2023)
 Calculate the bounds of your estimates if there is some violation

Most important advice:

- Do not think you are safe when event study does not detect anything
- Do power calculations against economically relevant violations of parallel trends (R package, but not Stata...)
- If you know the functional form of the differences in trends, control it
- Sensitivity analysis using Rambachan and Roth (2023)
 Calculate the bounds of your estimates if there is some violation

Most important advice:

- Do not think you are safe when event study does not detect anything
- Do power calculations against economically relevant violations of parallel trends (R package, but not Stata...)
- If you know the functional form of the differences in trends, control it
- Sensitivity analysis using Rambachan and Roth (2023)
 Calculate the bounds of your estimates if there is some violation

Most important advice:

- Do not think you are safe when event study does not detect anything
- Do power calculations against economically relevant violations of parallel trends (R package, but not Stata...)
- If you know the functional form of the differences in trends, control it
- Sensitivity analysis using Rambachan and Roth (2023)
 Calculate the bounds of your estimates if there is some violation

Most important advice:

- Do not think you are safe when event study does not detect anything
- Do power calculations against economically relevant violations of parallel trends (R package, but not Stata...)
- If you know the functional form of the differences in trends, control it
- Sensitivity analysis using Rambachan and Roth (2023)
 Calculate the bounds of your estimates if there is some violation

- The critical assumption for DID is parallel trend
- What if we do not have it?
- What if treated and control provinces have different trends?
- Let's create one control group! ⇒ Synthetic Control
- Synthetic control is a matching method

The critical assumption for DID is parallel trend

- What if we do not have it?
- What if treated and control provinces have different trends?
- Let's create one control group! \Rightarrow Synthetic Control
- Synthetic control is a matching method

The critical assumption for DID is parallel trend

What if we do not have it?

- What if treated and control provinces have different trends?
- Let's create one control group! \Rightarrow Synthetic Control
- Synthetic control is a matching method

- The critical assumption for DID is parallel trend
- What if we do not have it?
- What if treated and control provinces have different trends?
- Let's create one control group! \Rightarrow Synthetic Control
- Synthetic control is a matching method

- The critical assumption for DID is parallel trend
- What if we do not have it?
- What if treated and control provinces have different trends?
- Let's create one control group! \Rightarrow Synthetic Control
- Synthetic control is a matching method

- The critical assumption for DID is parallel trend
- What if we do not have it?
- What if treated and control provinces have different trends?
- Let's create one control group! \Rightarrow Synthetic Control
- Synthetic control is a matching method

All the following contexts come from Abadie, Diamond, and Hainmueller (2010, 2015); Abadie (2021)

The synthetic control method is based on the idea that, when the units of observation are a small number of aggregate entities, a combination of unaffected units often provides a more appropriate comparison than any single unaffected unit alone.

All the following contexts come from Abadie, Diamond, and Hainmueller (2010, 2015); Abadie (2021)

The synthetic control method is based on the idea that, when the units of observation are a small number of aggregate entities, a combination of unaffected units often provides a more appropriate comparison than any single unaffected unit alone. All the following contexts come from Abadie, Diamond, and Hainmueller (2010, 2015); Abadie (2021)

The synthetic control method is based on the idea that, when the units of observation are a small number of aggregate entities, a combination of unaffected units often provides a more appropriate comparison than any single unaffected unit alone.

- Take Abadie, Diamond, and Hainmueller (2010) as an example
- California implemented Proposition 99 in 1988
- It is a large-scale tobacco control program
 - A 25-cent per pack excise tax on the sale of tobacco cigarettes, cigars and chewing tobacco
 - A ban on cigarette vending machines in public areas accessible by juveniles

But it seems that pre-trends are very different across states



Figure 5. Per-capita cigarette sales gaps in California and placebo gaps in 34 control states (discards states with pre-Proposition 99 MSPE twenty times higher than California's).

▶ ≣|= ∽へ 36/52

-

• Even when you average over all control states, you have this


Synthetic Control: Main Idea

- Then you have to combine them to create a "synthetic" control state
- A man-made "synthetic" California



Figure 2. Trends in per-capita cigarette sales: California vs. syn thetic California.

Figure 3. Per-capita cigarette sales gap between California and synthetic California.

(a)

(b)

- Suppose we have j = 1, 2, ..., J + 1 units (provinces, cities...), spanning T periods
- **T**₀ is the treatment starting period, j = 1 is the treated unit
- We call j = 2, 3, ..., J + 1 as "donor pool"
- We will create the synthetic control group from units in this "donor pool"
- X_{kj} are observed characteristics, which can include pre-treatment values of Y
- X_{kj} are unaffected by treatment

- Suppose we have j = 1, 2, ..., J + 1 units (provinces, cities...), spanning T periods
- T_0 is the treatment starting period, j = 1 is the treated unit
- We call j = 2, 3, ..., J + 1 as "donor pool"
- We will create the synthetic control group from units in this "donor pool"
- X_{kj} are observed characteristics, which can include pre-treatment values of Y
- X_{kj} are unaffected by treatment

- Suppose we have j = 1, 2, ..., J + 1 units (provinces, cities...), spanning T periods
- T_0 is the treatment starting period, j = 1 is the treated unit
- We call j = 2, 3, ..., J + 1 as "donor pool"
- We will create the synthetic control group from units in this "donor pool"
- X_{kj} are observed characteristics, which can include pre-treatment values of Y
- X_{kj} are unaffected by treatment

- Suppose we have j = 1, 2, ..., J + 1 units (provinces, cities...), spanning T periods
- T_0 is the treatment starting period, j = 1 is the treated unit
- We call j = 2, 3, ..., J + 1 as "donor pool"
- We will create the synthetic control group from units in this "donor pool"
- X_{kj} are observed characteristics, which can include pre-treatment values of Y
- X_{kj} are unaffected by treatment

- Suppose we have j = 1, 2, ..., J + 1 units (provinces, cities...), spanning T periods
- T_0 is the treatment starting period, j = 1 is the treated unit
- We call j = 2, 3, ..., J + 1 as "donor pool"
- We will create the synthetic control group from units in this "donor pool"
- X_{kj} are observed characteristics, which can include pre-treatment values of Y
- X_{kj} are unaffected by treatment

- Suppose we have j = 1, 2, ..., J + 1 units (provinces, cities...), spanning T periods
- T_0 is the treatment starting period, j = 1 is the treated unit
- We call j = 2, 3, ..., J + 1 as "donor pool"
- We will create the synthetic control group from units in this "donor pool"
- X_{kj} are observed characteristics, which can include pre-treatment values of Y
- X_{kj} are unaffected by treatment

- Suppose we have j = 1, 2, ..., J + 1 units (provinces, cities...), spanning T periods
- T_0 is the treatment starting period, j = 1 is the treated unit
- We call j = 2, 3, ..., J + 1 as "donor pool"
- We will create the synthetic control group from units in this "donor pool"
- X_{kj} are observed characteristics, which can include pre-treatment values of Y
- X_{kj} are unaffected by treatment

• Define potential outcome: Y_{jt}^I, Y_{jt}^A

Treatment effect of interest:
$$\tau_{1t} = Y'_{jt} - Y'_{jt}$$
 for $t > T_0$

Treatment effect can vary across time

A synthetic control is defined as a weighted average of the units in the donor pool:

$$\hat{Y}_{1t}^N = \sum_{j=2}^{J+1} w_j Y_j$$

w_j is the weight assigned to donor j

Then we can estimate the treatment effect:

$$\tau_{1t} = Y_{1t} - \hat{Y}_{1t}^{\Lambda}$$

• Define potential outcome: Y'_{jt}, Y'_{jt}

Treatment effect of interest: $\tau_{1t} = Y_{jt}^{l} - Y_{jt}^{N}$ for $t > T_0$

Treatment effect can vary across time

A synthetic control is defined as a weighted average of the units in the donor pool:

$$\hat{Y}_{1t}^{N} = \sum_{j=2}^{J+1} w_j Y_{ji}$$

• w_j is the weight assigned to donor j

Then we can estimate the treatment effect:

$$\tau_{1t} = Y_{1t} - \hat{Y}_{1t}^{\Lambda}$$

《日》《問》《日》《日》 모님 《

• Define potential outcome: Y_{jt}^{I}, Y_{jt}^{N}

• Treatment effect of interest: $\tau_{1t} = Y'_{jt} - Y^N_{jt}$ for $t > T_0$

Treatment effect can vary across time

A synthetic control is defined as a weighted average of the units in the donor pool:

$$\hat{Y}_{1t}^N = \sum_{j=2}^{J+1} w_j Y_{ji}$$

• w_j is the weight assigned to donor j

Then we can estimate the treatment effect:

$$\tau_{1t} = Y_{1t} - \hat{Y}_{1t}^{\Lambda}$$

• Define potential outcome: Y_{jt}^{I}, Y_{jt}^{N}

• Treatment effect of interest: $\tau_{1t} = Y'_{jt} - Y^N_{jt}$ for $t > T_0$

Treatment effect can vary across time

A synthetic control is defined as a weighted average of the units in the donor pool:

$$\hat{Y}_{1t}^N = \sum_{j=2}^{J+1} w_j Y_{ji}$$

• w_j is the weight assigned to donor j

Then we can estimate the treatment effect:

$$\tau_{1t} = Y_{1t} - \hat{Y}_{1t}^{\Lambda}$$

《日》《問》《日》《日》 모님 《

• Define potential outcome: Y_{jt}^{I}, Y_{jt}^{N}

- Treatment effect of interest: $\tau_{1t} = Y'_{jt} Y^N_{jt}$ for $t > T_0$
- Treatment effect can vary across time
- A synthetic control is defined as a weighted average of the units in the donor pool:

$$\hat{Y}_{1t}^{N} = \sum_{j=2}^{J+1} w_j Y_{jt}$$

• w_j is the weight assigned to donor j

Then we can estimate the treatment effect:

$$\tau_{1t} = Y_{1t} - \hat{Y}_{1t}^{\Lambda}$$

《日》《問》《日》《日》 모님 《

• Define potential outcome: Y_{jt}^{I}, Y_{jt}^{N}

- Treatment effect of interest: $\tau_{1t} = Y'_{jt} Y^N_{jt}$ for $t > T_0$
- Treatment effect can vary across time
- A synthetic control is defined as a weighted average of the units in the donor pool:

$$\hat{Y}_{1t}^{N} = \sum_{j=2}^{J+1} w_j Y_{jt}$$

w_j is the weight assigned to donor j

Then we can estimate the treatment effect:

$$\tau_{1t} = Y_{1t} - \hat{Y}_{1t}^N$$

• Define potential outcome: Y_{jt}^{I}, Y_{jt}^{N}

- Treatment effect of interest: $\tau_{1t} = Y'_{jt} Y^N_{jt}$ for $t > T_0$
- Treatment effect can vary across time
- A synthetic control is defined as a weighted average of the units in the donor pool:

$$\hat{Y}_{1t}^N = \sum_{j=2}^{J+1} w_j Y_{jt}$$

- w_j is the weight assigned to donor j
- Then we can estimate the treatment effect:

$$\tau_{1t} = Y_{1t} - \hat{Y}_{1t}^{\Lambda}$$

・ロト・西ト・ヨト・ヨト 通道 の

- How to define the weights?
- We minimize the following:

$$||X_1 - X_0 W|| = \left(\sum_{h=1}^k v_h (X_{h1} - w_2 X_{h2} - \dots - w_{J+1} X_{hJ+1})^2\right)^{1/2}$$

- This is the weighted euclidean distance between X₁ and X₀
- We try to find a combination of donors that can mimic our treated group the best
- Watch out: the difference between weights v and weights w

How to define the weights?

• We minimize the following:

$$||X_1 - X_0 W|| = \left(\sum_{h=1}^k v_h (X_{h1} - w_2 X_{h2} - \dots - w_{J+1} X_{hJ+1})^2\right)^{1/2}$$

This is the weighted euclidean distance between X_1 and X_0

- We try to find a combination of donors that can mimic our treated group the best
- Watch out: the difference between weights *v* and weights *w*

- How to define the weights?
- We minimize the following:

$$||X_1 - X_0 W|| = \left(\sum_{h=1}^k v_h (X_{h1} - w_2 X_{h2} - \dots - w_{J+1} X_{hJ+1})^2\right)^{1/2}$$

- This is the weighted euclidean distance between X_1 and X_0
- We try to find a combination of donors that can mimic our treated group the best
- Watch out: the difference between weights *v* and weights *w*

- How to define the weights?
- We minimize the following:

$$||X_1 - X_0 W|| = \left(\sum_{h=1}^k v_h (X_{h1} - w_2 X_{h2} - \dots - w_{J+1} X_{hJ+1})^2\right)^{1/2}$$

- This is the weighted euclidean distance between X_1 and X_0
- We try to find a combination of donors that can mimic our treated group the best
- Watch out: the difference between weights v and weights w

- How to define the weights?
- We minimize the following:

$$||X_1 - X_0 W|| = \left(\sum_{h=1}^k v_h (X_{h1} - w_2 X_{h2} - \dots - w_{J+1} X_{hJ+1})^2\right)^{1/2}$$

- This is the weighted euclidean distance between X_1 and X_0
- We try to find a combination of donors that can mimic our treated group the best
- Watch out: the difference between weights v and weights w

- How to define the weights?
- We minimize the following:

$$||X_1 - X_0 W|| = \left(\sum_{h=1}^k v_h (X_{h1} - w_2 X_{h2} - \dots - w_{J+1} X_{hJ+1})^2\right)^{1/2}$$

- This is the weighted euclidean distance between X_1 and X_0
- We try to find a combination of donors that can mimic our treated group the best
- Watch out: the difference between weights v and weights w

- w is the weight assigned to each unit (state) when we want to create a synthetic control group (state)
- v is the weight assigned to each characteristic when we try to calculate w

- w is the weight assigned to each unit (state) when we want to create a synthetic control group (state)
- \bullet v is the weight assigned to each characteristic when we try to calculate w

- w is the weight assigned to each unit (state) when we want to create a synthetic control group (state)
- v is the weight assigned to each characteristic when we try to calculate w

How to estimate the effect of the 1990 German reunification
 Treated: West Germany; Untreated: Other OECD countries

How to estimate the effect of the 1990 German reunification Treated: West Germany; Untreated: Other OECD countries

How to estimate the effect of the 1990 German reunificationTreated: West Germany; Untreated: Other OECD countries

v are weights for economic predictors: The importance of each predictor for the match procedure

Ec	TABLE 1 Economic Crowth Predictor Means before the German Reunification				
	West Germany (1)	Synthetic West Germany (2)	OECD average (3)	Austria (nearest neighbor) (4)	
GDP per capita	15,808.9	15,802.2	13,669.4	14,817.0	
Trade openness	56.8	56.9	59.8	74.6	
Inflation rate	2.6	3.5	7.6	3.5	
Industry share	34.5	34.4	33.8	35.5	
Schooling	55.5	55.2	38.7	60.9	
Investment rate	27.0	27.0	25.9	26.6	

• *w* are weights for compared countries: the importance of each country in forming the synthetic Germany

TABLE 2 Synthetic Control Weights	for West Germany
Australia	
Austria	0.42
Belgium	
Denmark	_
France	_
Greece	
Italy	
Japan	0.16
Netherlands	0.09
New Zealand	
Norway	_
Portugal	_
Spain	
Switzerland	0.11
United Kingdom	_
United States	0.22

Synthetic Control: Procedure

How to determine v?

- Step 1: Divide all pre-treatment sample into 2 parts Part 1 $t = 1, ..., t_0$ and Part 2 $t = t_0 + 1, ..., T_0$
- Step 2: Find the best V^{*} that minimizes MSPE for Part 1 data:

$$\sum_{t=t_0+1}^{T_0} (Y_{1t} - w_2(V)Y_{2t} - \dots - w_{J+1}(V)Y_{J+1t})^2$$

w(V) are the unit weights given each value of predictor weights V
 Step 3: Using V^{*} and Part 2 data to calculate W

Synthetic Control: Procedure

How to determine v?

- Step 1: Divide all pre-treatment sample into 2 parts
- Part 1 $t = 1, ..., t_0$ and Part 2 $t = t_0 + 1, ..., T_0$
- Step 2: Find the best V^{*} that minimizes MSPE for Part 1 data:

$$\sum_{t=t_0+1}^{T_0} (Y_{1t} - w_2(V)Y_{2t} - \dots - w_{J+1}(V)Y_{J+1t})^2$$

w(V) are the unit weights given each value of predictor weights V
 Step 3: Using V^{*} and Part 2 data to calculate W

Synthetic Control: Procedure

How to determine v?

- Step 1: Divide all pre-treatment sample into 2 parts Part 1 $t = 1, ..., t_0$ and Part 2 $t = t_0 + 1, ..., T_0$
- Step 2: Find the best V^{*} that minimizes MSPE for Part 1 data:

$$\sum_{t=t_0+1}^{T_0} (Y_{1t} - w_2(V)Y_{2t} - \dots - w_{J+1}(V)Y_{J+1t})^2$$

w(V) are the unit weights given each value of predictor weights V
Step 3: Using V* and Part 2 data to calculate W

How to determine v?

- Step 1: Divide all pre-treatment sample into 2 parts Part 1 $t = 1, ..., t_0$ and Part 2 $t = t_0 + 1, ..., T_0$
- Step 2: Find the best V* that minimizes MSPE for Part 1 data:

$$\sum_{t=t_0+1}^{T_0} (Y_{1t} - w_2(V)Y_{2t} - \dots - w_{J+1}(V)Y_{J+1t})^2$$

w(V) are the unit weights given each value of predictor weights V
Step 3: Using V* and Part 2 data to calculate W

How to determine v?

- Step 1: Divide all pre-treatment sample into 2 parts Part 1 $t = 1, ..., t_0$ and Part 2 $t = t_0 + 1, ..., T_0$
- Step 2: Find the best V* that minimizes MSPE for Part 1 data:

$$\sum_{t=t_0+1}^{T_0} (Y_{1t} - w_2(V)Y_{2t} - \dots - w_{J+1}(V)Y_{J+1t})^2$$

w(V) are the unit weights given each value of predictor weights V
Step 3: Using V* and Part 2 data to calculate W

Synthetic West Germany





Homework: Explain the reason why we split the data into two parts. No math!
- Post-treatment outcomes cannot be used in X! (but pre-treatment outcomes can)
- There can be bias if there are unobserved endogenous characteristics that cannot be matched
- When the difference $X_1 X_0 W^*$ is large (pre-treatment fit is bad), do not use synthetic control
- If the size of the donor pool is too large, there can be over-fitting
- So, choose the donor pool judiciously
- All weights must be within [0,1]

- Post-treatment outcomes cannot be used in X! (but pre-treatment outcomes can)
- There can be bias if there are unobserved endogenous characteristics that cannot be matched
- When the difference X₁ X₀W^{*} is large (pre-treatment fit is bad), do not use synthetic control
- If the size of the donor pool is too large, there can be over-fitting
- So, choose the donor pool judiciously
- All weights must be within [0,1]

- Post-treatment outcomes cannot be used in X! (but pre-treatment outcomes can)
- There can be bias if there are unobserved endogenous characteristics that cannot be matched
- When the difference $X_1 X_0 W^*$ is large (pre-treatment fit is bad), do not use synthetic control
- If the size of the donor pool is too large, there can be over-fitting
- So, choose the donor pool judiciously
- All weights must be within [0,1]

- Post-treatment outcomes cannot be used in X! (but pre-treatment outcomes can)
- There can be bias if there are unobserved endogenous characteristics that cannot be matched
- When the difference $X_1 X_0 W^*$ is large (pre-treatment fit is bad), do not use synthetic control
- If the size of the donor pool is too large, there can be over-fitting
- So, choose the donor pool judiciously
- All weights must be within [0,1]

- Post-treatment outcomes cannot be used in X! (but pre-treatment outcomes can)
- There can be bias if there are unobserved endogenous characteristics that cannot be matched
- When the difference $X_1 X_0 W^*$ is large (pre-treatment fit is bad), do not use synthetic control
- If the size of the donor pool is too large, there can be over-fitting
- So, choose the donor pool judiciously
- All weights must be within [0,1]

- Post-treatment outcomes cannot be used in X! (but pre-treatment outcomes can)
- There can be bias if there are unobserved endogenous characteristics that cannot be matched
- When the difference $X_1 X_0 W^*$ is large (pre-treatment fit is bad), do not use synthetic control
- If the size of the donor pool is too large, there can be over-fitting
- So, choose the donor pool judiciously
- All weights must be within [0,1]

- Post-treatment outcomes cannot be used in X! (but pre-treatment outcomes can)
- There can be bias if there are unobserved endogenous characteristics that cannot be matched
- When the difference $X_1 X_0 W^*$ is large (pre-treatment fit is bad), do not use synthetic control
- If the size of the donor pool is too large, there can be over-fitting
- So, choose the donor pool judiciously
- All weights must be within [0,1]

- Post-treatment outcomes cannot be used in X! (but pre-treatment outcomes can)
- There can be bias if there are unobserved endogenous characteristics that cannot be matched
- When the difference $X_1 X_0 W^*$ is large (pre-treatment fit is bad), do not use synthetic control
- If the size of the donor pool is too large, there can be over-fitting
- So, choose the donor pool judiciously
- All weights must be within [0,1]

- When we have data across time for different units, we have panel data
- FE, dummy, and FD regressions can cancel out time-invariant confounders
- FE and dummy regressions are identical
- FE and FD are identical in 2-period cases, but different for more than 2
- FE is generally more efficient and thus, preferred

When we have data across time for different units, we have panel data

- FE, dummy, and FD regressions can cancel out time-invariant confounders
- FE and dummy regressions are identical
- **FE** and FD are identical in 2-period cases, but different for more than 2
- FE is generally more efficient and thus, preferred

- When we have data across time for different units, we have panel data
- FE, dummy, and FD regressions can cancel out time-invariant confounders
- FE and dummy regressions are identical
- **FE** and FD are identical in 2-period cases, but different for more than 2
- FE is generally more efficient and thus, preferred

- When we have data across time for different units, we have panel data
- FE, dummy, and FD regressions can cancel out time-invariant confounders
- FE and dummy regressions are identical
- **FE** and FD are identical in 2-period cases, but different for more than 2
- FE is generally more efficient and thus, preferred

- When we have data across time for different units, we have panel data
- FE, dummy, and FD regressions can cancel out time-invariant confounders
- FE and dummy regressions are identical
- FE and FD are identical in 2-period cases, but different for more than 2
- FE is generally more efficient and thus, preferred

- When we have data across time for different units, we have panel data
- FE, dummy, and FD regressions can cancel out time-invariant confounders
- FE and dummy regressions are identical
- FE and FD are identical in 2-period cases, but different for more than 2
- FE is generally more efficient and thus, preferred

- The main assumption for DID is parallel trend
- Traditionally, we validate parallel trend assumption in two steps:
 - Draw a figure
 - Run event study
- However, statistical test minimizes T1ER, which inflates T2ER
- Be careful using them!
- Check the power and validate your assumption using economic knowledge

The main assumption for DID is parallel trend

Traditionally, we validate parallel trend assumption in two steps:

- Draw a figure
- Run event study
- However, statistical test minimizes T1ER, which inflates T2ER
- Be careful using them!
- Check the power and validate your assumption using economic knowledge

- The main assumption for DID is parallel trend
- Traditionally, we validate parallel trend assumption in two steps:
 - Draw a figure
 - Run event study
- However, statistical test minimizes T1ER, which inflates T2ER
- Be careful using them!
- Check the power and validate your assumption using economic knowledge

- The main assumption for DID is parallel trend
- Traditionally, we validate parallel trend assumption in two steps:
 - Draw a figure
 - Run event study
- However, statistical test minimizes T1ER, which inflates T2ER
- Be careful using them!
- Check the power and validate your assumption using economic knowledge

- The main assumption for DID is parallel trend
- Traditionally, we validate parallel trend assumption in two steps:
 - Draw a figure
 - Run event study
- However, statistical test minimizes T1ER, which inflates T2ER
- Be careful using them!
- Check the power and validate your assumption using economic knowledge

- The main assumption for DID is parallel trend
- Traditionally, we validate parallel trend assumption in two steps:
 - Draw a figure
 - Run event study
- However, statistical test minimizes T1ER, which inflates T2ER
- Be careful using them!
- Check the power and validate your assumption using economic knowledge

- The main assumption for DID is parallel trend
- Traditionally, we validate parallel trend assumption in two steps:
 - Draw a figure
 - Run event study
- However, statistical test minimizes T1ER, which inflates T2ER
- Be careful using them!
- Check the power and validate your assumption using economic knowledge

- The main assumption for DID is parallel trend
- Traditionally, we validate parallel trend assumption in two steps:
 - Draw a figure
 - Run event study
- However, statistical test minimizes T1ER, which inflates T2ER
- Be careful using them!
- Check the power and validate your assumption using economic knowledge

- When it is hard to find a good control group with parallel trend
- You can create one using synthetic control method
- You assign two sets of weights and taking the weighted average
 - Weights for each characteristics.
 - Weights for each donor unit
- Then you create a control group to mimic the behavior of the treated group

• When it is hard to find a good control group with parallel trend

- Vou can create one using synthetic control method
- You assign two sets of weights and taking the weighted average
 - Weights for each characteristics
 - Weights for each donor unit

Then you create a control group to mimic the behavior of the treated group

• When it is hard to find a good control group with parallel trend

You can create one using synthetic control method

- You assign two sets of weights and taking the weighted average
 - Weights for each characteristics
 - Weights for each donor unit
- Then you create a control group to mimic the behavior of the treated group

- When it is hard to find a good control group with parallel trend
- You can create one using synthetic control method
- You assign two sets of weights and taking the weighted average
 - Weights for each characteristics
 - Weights for each donor unit
- Then you create a control group to mimic the behavior of the treated group

- When it is hard to find a good control group with parallel trend
- You can create one using synthetic control method
- You assign two sets of weights and taking the weighted average
 - Weights for each characteristics
 - Weights for each donor unit

Then you create a control group to mimic the behavior of the treated group

- When it is hard to find a good control group with parallel trend
- You can create one using synthetic control method
- You assign two sets of weights and taking the weighted average
 - Weights for each characteristics
 - Weights for each donor unit

Then you create a control group to mimic the behavior of the treated group

- When it is hard to find a good control group with parallel trend
- You can create one using synthetic control method
- You assign two sets of weights and taking the weighted average
 - Weights for each characteristics
 - Weights for each donor unit
- Then you create a control group to mimic the behavior of the treated group

- Abadie, Alberto. 2021. "Using Synthetic Controls: Feasibility, Data Requirements, and Methodological Aspects." *Journal of Economic Literature* 59 (2):391–425.
- Abadie, Alberto, Alexis Diamond, and Jens Hainmueller. 2010. "Synthetic Control Methods for Comparative Case Studies: Estimating the Effect of California's Tobacco Control Program." *Journal of the American Statistical Association* 105 (490):493–505.
- ------. 2015. "Comparative Politics and the Synthetic Control Method." *American Journal of Political Science* 59 (2):495–510.
- Card, David and Alan B Krueger. 1994. "Minimum Wages and Employment: A Case Study of the Fast-Food Industry in New Jersey and Pennsylvania." *The American Economic Review* 84 (4):772–793.
- Rambachan, Ashesh and Jonathan Roth. 2023. "A More Credible Approach to Parallel Trends." *Review of Economic Studies* :rdad018.
- Roth, Jonathan. 2022. "Pretest with Caution: Event-study Estimates after Testing for Parallel Trends." American Economic Review: Insights 4 (3):305–322.